

# Transport Reversible Jump Proposals

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## Motivation

The problem of interest is sampling probability distribution  $\pi$  on

$$\mathcal{X} = \bigcup_{k \in \mathcal{K}} (\{k\} \times \Theta_k), \quad (1)$$

with parameters  $\theta_k \in \Theta_k \subseteq \mathbb{R}^{n_k}$  and model index (or indicator)  $k \in \mathcal{K}$ . We want to make inference on the joint distribution (or conditional factorization)  $\pi(k, \theta_k) = \pi(k)\pi(\theta_k|k)$ . When data  $\mathbf{y}$  is introduced, this becomes  $\pi(k, \theta_k|\mathbf{y}) = \pi(k|\mathbf{y})\pi(\theta_k|k, \mathbf{y})$ .

## Contributions

- I. A new class of RJMCMC proposals, called *transport reversible jump* (TRJ) proposals are developed that have desirable properties (see Proposition 1).
- II. Efficacy of the proposed approach is demonstrated in numerical studies on challenging examples using *approximate* transport maps.
- III. A modified version of the model probability estimator of [1] is applied to assess the quality proposals.
- IV. An alternative “all-in-one” approach to training approximate TMs using *conditional normalizing flows* is explored (see paper).

## Reversible Jump Markov Chain Monte Carlo

We want to propose from point  $\mathbf{x}$  to point  $\mathbf{x}'$ , noting  $\theta_k, \theta_{k'}$  have dimensions  $n_k, n_{k'}$  respectively. Following [2], we:

- Require dimensions match: introduce auxiliary variables  $\mathbf{u}_k \in \mathcal{U}_{k,k'} \subseteq \mathbb{R}^{w_k}$  and  $\mathbf{u}_{k'} \in \mathcal{U}_{k',k} \subseteq \mathbb{R}^{w_{k'}}$  such that  $n_k + w_k = n_{k'} + w_{k'}$ .
- Choose a diffeomorphism  $e$ .  $\theta_{k'}, \mathbf{u}_{k'} = h_{k,k'}(\theta_k, \mathbf{u}_k)$ .

A (simplified) RJMCMC Algorithm when  $n_{k'} > n_k$  is:

1. Propose model index  $k' \sim j_k(\cdot)$
2. Propose auxiliary variables  $\mathbf{u}_{k'} \sim g_{k,k'}(\cdot)$
3. Accept with probability

$$\alpha(\mathbf{x}, \mathbf{x}') = 1 \wedge \frac{\pi(\mathbf{x}')j_{k'}(k)g_{k',k}(\mathbf{u}_{k'})}{\pi(\mathbf{x})j_k(k')g_{k,k'}(\mathbf{u}_k)} \bigg|_{J_{h_{k,k'}}(\theta_k, \mathbf{u}_k)}. \quad (2)$$

## Transport Maps and Normalizing Flows

A function  $T: \mathbb{R}^n \rightarrow \mathbb{R}^n$  is called a *transport map* (TM) from distribution  $\mu_\theta$  to distribution  $\mu_{\mathbf{z}}$  if  $\mu_{\mathbf{z}} = T_{\#}\mu_\theta$ , i.e.  $\mu_{\mathbf{z}}$  is the *pushforward* of  $\mu_\theta$  using the measurable function  $T$ .

Let  $\{T_\psi\}$  be a family of *diffeomorphisms* with domain on the support of some arbitrary base distribution  $\mu_{\mathbf{z}}$ . Then, for fixed parameters  $\psi$ , the PDF of the random vector  $\vartheta = T_\psi(\mathbf{z})$  is

$$\mu_\vartheta(\vartheta; \psi) = \mu_{\mathbf{z}}(T_\psi^{-1}(\vartheta)) |J_{T_\psi^{-1}}(\vartheta)|, \quad \vartheta \in \mathbb{R}^n. \quad (3)$$

Distributions  $\mu_\vartheta$  are *flow-based models*, where  $\{T_\psi\}$  are the *normalizing flows* [5]. With finite samples  $\mathbf{s} \sim \pi$ , we obtain an *approximate* TM  $\hat{T}$  via density estimation, minimising the KLD from  $\{\mathbf{s}\}$  to  $\mu_\vartheta$ .

## Proposed Method: Transport Reversible Jump

First, let  $\nu$  be a univariate closed-form distribution e.g. standard normal  $\phi$ . For each model  $k \in \mathcal{M}$ :

- Define reference distributions  $\otimes_{n_k}\nu$ .
- Train flows on samples  $\mathbf{s}_k \sim \pi(\theta_k|k)$  to obtain  $\hat{T}_k$ .

Then, a transdimensional proposal where  $n_{k'} > n_k$  is

$$\begin{aligned} \mathbf{z}_k &\leftarrow T_k(\theta_k), \\ \mathbf{z}'_{k'} &\leftarrow \bar{h}_{k,k'}(\mathbf{z}_k, \mathbf{u}_k), \\ \theta_{k'} &\leftarrow T_{k'}^{-1}(\mathbf{z}'_{k'}), \end{aligned} \quad (4)$$

where  $\bar{h}_{k,k'}$  is a *volume-preserving* diffeomorphism on  $\otimes_{n_k}\nu$ .

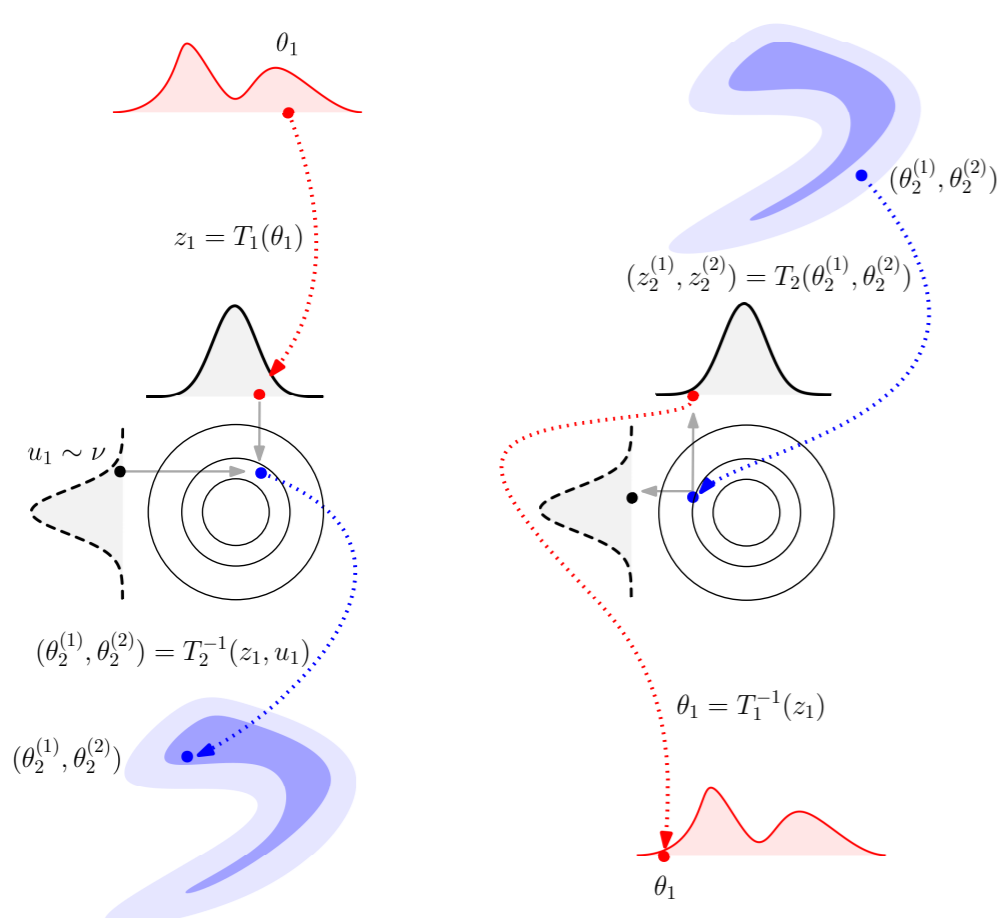


Figure 1. Illustration of the proposal class. Here, the reference  $\nu$  is Gaussian. The diffeomorphisms  $(\bar{h}_{k,k'})$  on the reference distributions concatenate or extract coordinates as required.

## Proposition: Exact Transport RJMCMC

Suppose that RJMCMC proposals are of the form described in (4), and for each  $k \in \mathcal{K}$ , satisfy  $T_k \# \pi_k = \otimes_{n_k}\nu$ . Then, (2) reduces to

$$\alpha(\mathbf{x}, \mathbf{x}') = 1 \wedge \frac{\pi(k')j_{k'}(k)}{\pi(k)j_k(k')}. \quad (5)$$

## Corollary

Provided the conditions of Proposition 1 are satisfied, choosing  $\{j_k\}$  such that

$$\pi(k')j_{k'}(k) = \pi(k)j_k(k'), \quad \forall k, k' \in \mathcal{K}, \quad (6)$$

leads to a rejection-free proposal.

## Illustrative Example: 1D 2D Sinh-Arcsinh Target

Using the inverse sinh-arcsinh transformation of [3]

$$S_{\epsilon, \delta}(\cdot) = \sinh(\delta^{-1} \odot (\sinh^{-1}(\cdot) + \epsilon)), \quad \epsilon \in \mathbb{R}^n, \delta \in \mathbb{R}_+^n.$$

For  $\mathbf{Z} \sim \mathcal{N}(\mathbf{0}_n, \mathbf{I}_{n \times n})$  and lower triangular  $n \times n$  matrix  $L$ , the exact (or “perfect”) transport is

$$T(\mathbf{Z}) = S_{\epsilon, \delta}(L\mathbf{Z}), \text{ i.e. } T^{-1}(\cdot) = L^{-1}S_{\epsilon, \delta}^{-1}(\cdot), \quad (7)$$

for chosen reference distributions  $\phi_n, n_k = k$ . The PDF for  $\theta = T(\mathbf{Z})$  takes the form in (3). The target of interest for this example, where  $\theta_1 = (\theta_1^{(1)})$  and  $\theta_2 = (\theta_2^{(1)}, \theta_2^{(2)})$ , is

$$\pi(k, \theta_k) = \begin{cases} \frac{1}{4} p_{\epsilon_1, \delta_1, 1}(\theta_1), & k = 1, \\ \frac{3}{4} p_{\epsilon_2, \delta_2, L}(\theta_2), & k = 2, \end{cases} \quad (8)$$

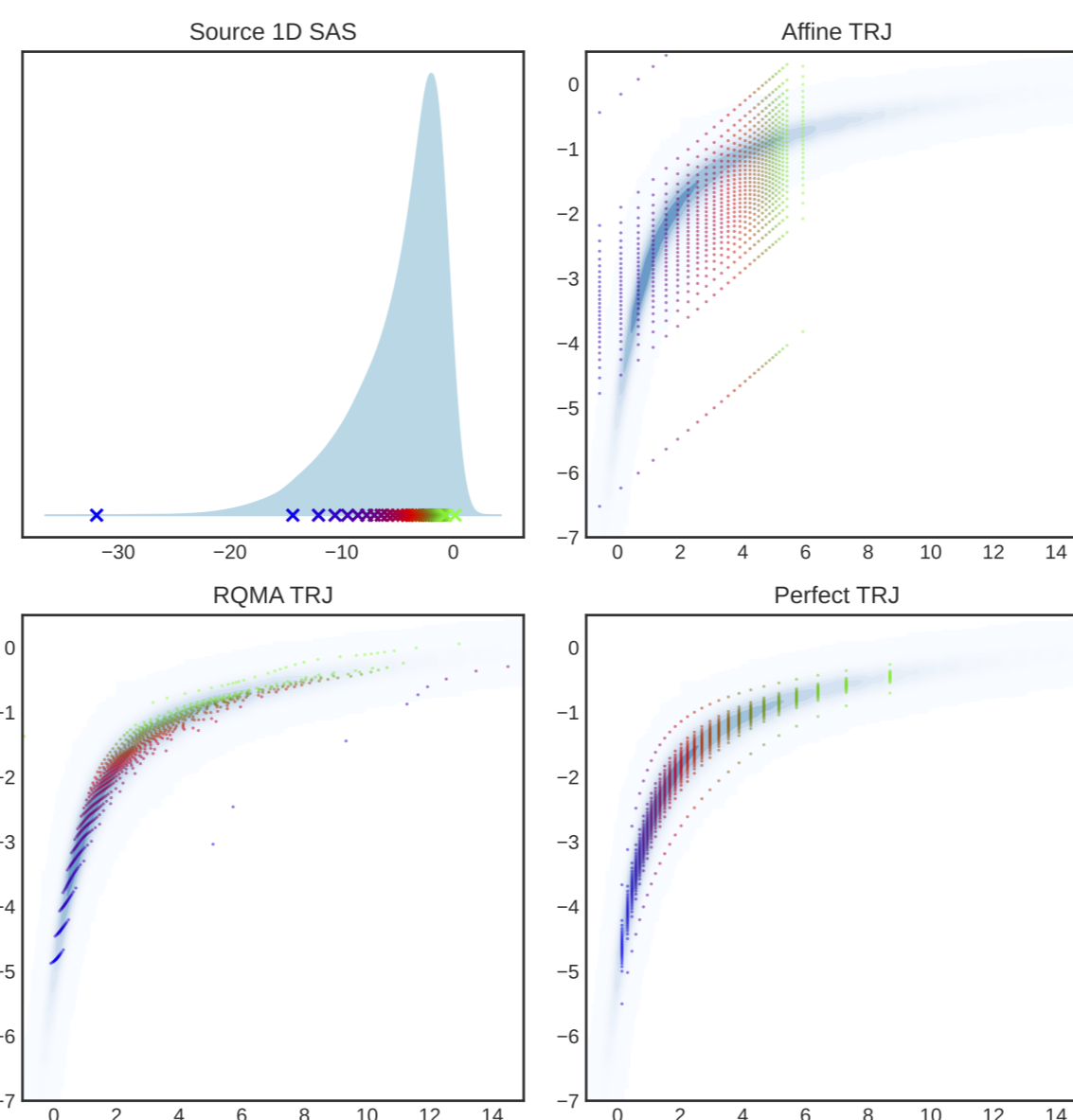


Figure 2. Systematic draws from conditional target  $\pi(x_1|k=1)$  of (8) are transported from  $(1, \theta_1) \in \mathcal{K} \times \mathbb{R}^1$  (top left) to  $(2, (\theta_1, \theta_2)) \in \mathcal{K} \times \mathbb{R}^2$  via TRJ proposals using: Top right Approximate affine, Bottom left Approximate RQMA-NF, Bottom right Perfect TM. The auxiliary variables in the proposals are also drawn systematically (30 for each point in the source distribution).

## Modified Bartolucci Bridge Sampling Estimator

For an RJMCMC chain, [1] showed that the Bayes factor  $B_{k,k'}$  (ratio of marginal likelihoods) is estimated via

$$\hat{B}_{k,k'} = \frac{N_{k'}^{-1} \sum_{i=1}^{N_{k'}} \alpha'_i}{N_k^{-1} \sum_{i=1}^{N_k} \alpha_i}, \quad (9)$$

where  $N_{k'}$  and  $N_k$  are the number of proposed moves from model  $k'$  to  $k$ , and from  $k$  to  $k'$ , respectively. Assuming uniform prior model probabilities, estimates of posterior model probabilities are

$$\hat{\pi}(k) = \hat{B}_{j,k}^{-1} \left( 1 + \sum_{i \in \mathcal{K} \setminus \{j\}} \hat{B}_{i,j} \right)^{-1}, \text{ for arbitrary } j \in \mathcal{K}. \quad (10)$$

The **Modified Bartolucci Estimator** (MBE) adopts the above for proposals from *samples* of the conditional targets.

## References

- [1] Francesco Bartolucci, Luisa Scaccia, and Antonietta Mira. Efficient Bayes factor estimation from the reversible jump output. *Biometrika*, 93(1):41–52, 2006.
- [2] Peter J Green. Reversible jump Markov chain Monte Carlo computation and Bayesian model determination. *Biometrika*, 82(4):711–732, 1995.
- [3] M. C. Jones and Arthur Pewsey. Sinh-arcsinh distributions. *Biometrika*, 96(4):761–780, December 2009.
- [4] Hedibert Freitas Lopes and Mike West. Bayesian model assessment in factor analysis. *Statistica Sinica*, 14(1):41–67, 2004.
- [5] Danilo Rezende and Shakir Mohamed. Variational inference with normalizing flows. In *Proceedings of the 32nd International Conference on Machine Learning*, pages 1530–1538. PMLR, June 2015.

## Bayesian Factor Analysis Example

Task is to select a model for monthly exchange rates of six currencies relative to the British pound, spanning January 1975 to December 1986, denoted as  $\mathbf{y}_i \in \mathbb{R}^6$  for  $i = 1, \dots, 143$ , of the random vector  $\mathbf{Y}$ . Assume  $\mathbf{Y} \sim \mathcal{N}(\mathbf{0}_6, \Sigma)$ , where  $k$  is no. factors,  $\Sigma = \beta_k \beta_k^\top + \Lambda$ ,  $\Lambda$  is a  $6 \times 6$  positive diagonal matrix, and  $\beta_k$  is a  $6 \times k$  lower-triangular matrix with a positive diagonal. Following [4], for each  $\beta_k = [\beta_{ij}]$  with  $i = 1, \dots, 6, j = 1, \dots, k$ , the priors are

$$\begin{aligned} \beta_{ij} &\sim \mathcal{N}(0, 1), \quad i < j \\ \beta_{ii} &\sim \mathcal{N}_+(0, 1), \\ \Lambda_{ii} &\sim \mathcal{IG}(1.1, 0.05), \end{aligned} \quad (11)$$

Posterior probability of  $\theta_k = (\beta_k, \Lambda)$  for  $k = 2, 3$  via Bayes' Thm

$$\pi(k, \theta_k|\mathbf{y}) \propto p(k)p(\beta_k|k)p(\Lambda) \prod_{i=1}^{143} \phi_{\beta\beta^\top + \Lambda}(\mathbf{y}_i). \quad (12)$$

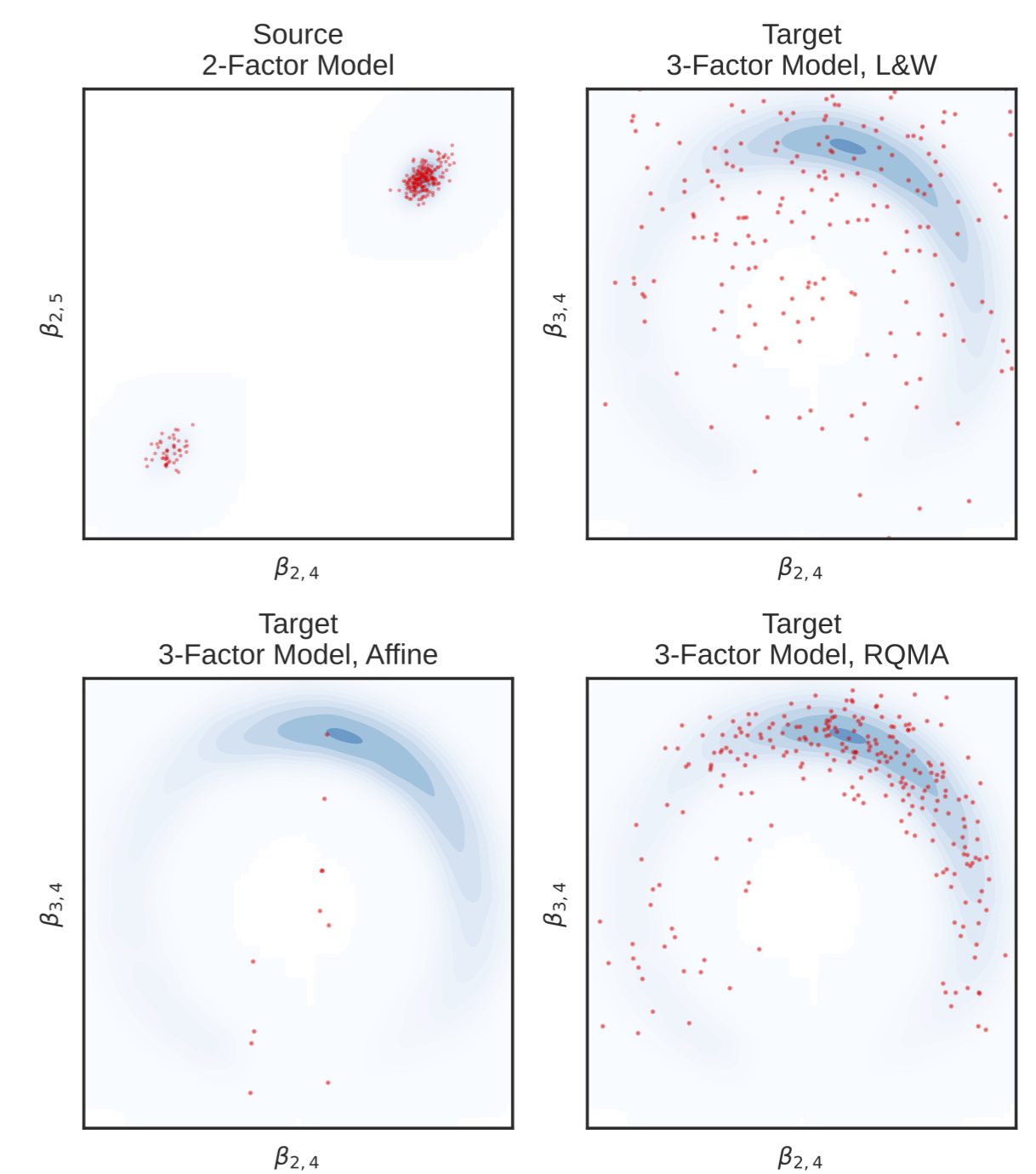


Figure 3. A visualization (using selected bivariate plots) of the proposal from points on the 2-factor model (top-left) to proposed points on the 3-factor model for each proposal type: [4] independence RJMCMC proposal (top right); TRJ with Affine map (bottom left); TRJ with RQMA-NF map (bottom right).

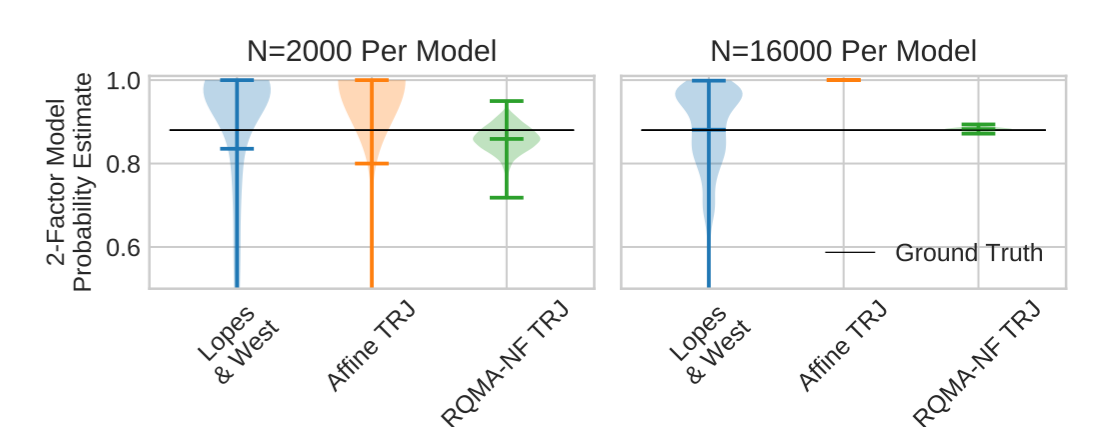


Figure 4. Violin plot showing the variability of the 2-factor model probability estimates in the case where only the 2-factor and 3-factor models are compared. Model probability estimates are obtained via the MBE. Ground truth is estimated via extended individual SMC runs ( $N = 5 \cdot 10^4$ ).

## Example: Robust Regression Variable Selection

Please see the paper (QR code below) for a detailed description.

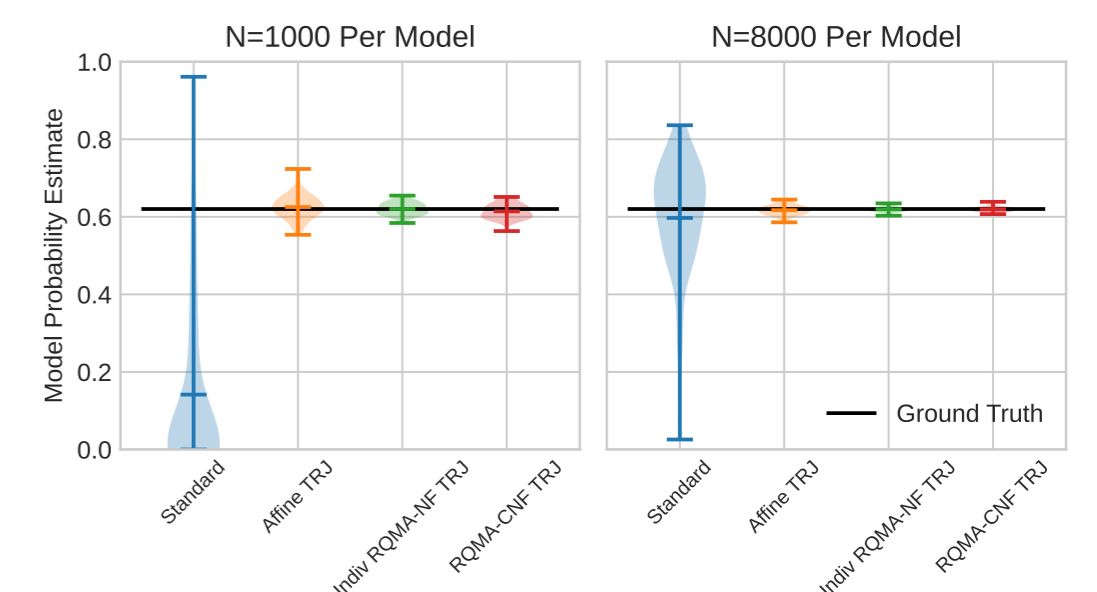


Figure 5. Violin plot showing the variability of the  $k = (1, 1, 1, 1)$  model probability estimate for each proposal type using the MBE vs ground truth individual SMC ( $N = 5 \cdot 10^4$ ). Individual SMC with  $N = 1000, 8000$  particles sampled conditional targets split into training/test samples for a total of 80 passes.

## Conclusions and Future Research Avenues

- Theoretically rejection-free approach for RJMCMC.
- Proposals using approximate TMs outperform baseline.
- Caveat of requiring samples to train the approximate TMs.
- Conditional NF approach to reduce training time (see paper).
- Efforts are justified in expensive-likelihood scenarios.
- MBE benchmark performance justifies further investigation.

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